$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$

$$\frac{\frac{d(e^{cx})}{dx}}{\frac{d(x^2)}{dx}} = ce^{cx}.$$

$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$
$$\frac{d(x^2)}{dx} = 2x.$$
$$\int x dx = \frac{x^2}{2} + c.$$

$$\frac{d(e^{cx})}{dx} = Ce^{cx}.$$
$$\frac{d(x^2)}{dx} = 2x.$$
$$\int x dx = \frac{x^2}{2} + c.$$
$$\frac{d(\ln x)}{dx} = \frac{1}{x}$$

$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$

$$\frac{d(x^2)}{dx} = 2x.$$

$$\int x dx = \frac{x^2}{2} + c.$$

$$\frac{d(\ln x)}{dx} = \frac{1}{x}$$
Chain Rule: $\frac{d(f(g(x)))}{dx} = f'(g(x))g'(x)dx$

$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$

$$\frac{d(x^2)}{dx} = 2x.$$

$$\int x dx = \frac{x^2}{2} + c.$$

$$\frac{d(\ln x)}{dx} = \frac{1}{x}$$
Chain Rule: $\frac{d(f(g(x)))}{dx} = f'(g(x))g'(x)dx$
Product Rule:

$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$

$$\frac{d(x^2)}{dx} = 2x.$$

$$\int x dx = \frac{x^2}{2} + c.$$

$$\frac{d(\ln x)}{dx} = \frac{1}{x}$$

Chain Rule: $\frac{d(f(g(x)))}{dx} = f'(g(x))g'(x)dx$

Product Rule:

(f(x)g(x))' = f'(x)g(x) + f(x)g'(x).

$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$
$$\frac{d(x^2)}{dx} = 2x.$$
$$\int x dx = \frac{x^2}{2} + c.$$
$$\frac{d(\ln x)}{dx} = \frac{1}{x}$$

Chain Rule: $\frac{d(f(g(x)))}{dx} = f'(g(x))g'(x)dx$

Product Rule:

(f(x)g(x))' = f'(x)g(x) + f(x)g'(x).d(uv) = udv + vdu

$$\frac{d(e^{cx})}{dx} = ce^{cx}.$$
$$\frac{d(x^2)}{dx} = 2x.$$
$$\int x dx = \frac{x^2}{2} + c.$$
$$\frac{d(\ln x)}{dx} = \frac{1}{x}$$

Chain Rule: $\frac{d(f(g(x)))}{dx} = f'(g(x))g'(x)dx$

Product Rule:

$$(f(x)g(x))' = f'(x)g(x) + f(x)g'(x).$$

$$d(uv) = udv + vdu$$

Integration by Parts: $\int u dv = uv - \int v du$.



Continuous Probability 1

1. pdf:



Continuous Probability 1

1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$.



Continuous Probability 1

1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$.

2. CDF:

Continuous Probability 1

1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$.

2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y) dy$.

- 1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$.
- 2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i) \delta = \int_{-\infty}^x f_X(y) dy$.
- 3. *X* ~ *U*[*a*,*b*]:

Continuous Probability 1

1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$. 2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$. 3. $X \sim U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\};$

Continuous Probability 1

1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$. 2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$. 3. $X \sim U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a}$ for $a \le x \le b$.

Continuous Probability 1

1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$. 2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$. 3. $X \sim U[a, b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a} \text{ for } a \le x \le b$. 4. $X \sim Expo(\lambda)$:

1. pdf:
$$Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$$
.
2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$.
3. $X \sim U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a} \text{ for } a \le x \le b$.
4. $X \sim Expo(\lambda)$:
 $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \ge 0\};$

1. pdf:
$$Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$$
.
2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$.
3. $X \sim U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a} \text{ for } a \le x \le b$.
4. $X \sim Expo(\lambda)$:
 $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \ge 0\}; F_X(x) = 1 - \exp\{-\lambda x\} \text{ for } x \le 0$.

1. pdf:
$$Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$$
.
2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$.
3. $X \sim U[a, b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a} \text{ for } a \le x \le b$.
4. $X \sim Expo(\lambda)$:
 $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \ge 0\}; F_X(x) = 1 - \exp\{-\lambda x\} \text{ for } x \le 0$.
5. Target:

1. pdf:
$$Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$$
.
2. CDF: $Pr[X \le x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^x f_X(y)dy$.
3. $X \sim U[a, b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a}$ for $a \le x \le b$.
4. $X \sim Expo(\lambda)$:
 $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \ge 0\}; F_X(x) = 1 - \exp\{-\lambda x\}$ for $x \le 0$.
5. Target: $f_X(x) = 2x \cdot 1\{0 \le x \le 1\};$

- 1. pdf: $Pr[X \in (x, x + \delta]] \approx f_X(x)\delta$. 2. CDF: $Pr[X < x] = F_X(x) = \lim_{\delta \to 0} \sum_i f_X(x_i) \delta = \int_{-\infty}^x f_X(y) dy$. 3. $X \sim U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}$; $F_X(x) = \frac{x-a}{b-a}$ for $a \le x \le b$. 4. $X \sim Expo(\lambda)$: $f_X(x) = \lambda \exp\{-\lambda x\} \mathbf{1}\{x \ge 0\}; F_X(x) = 1 - \exp\{-\lambda x\} \text{ for } x \le 0.$ 5. Target: $f_X(x) = 2x \cdot 1\{0 \le x \le 1\}$; $F_X(x) = x^2$ for $0 \le x \le 1$. 6. Joint pdf: $Pr[X \in (x, x + \delta), Y = (y, y + \delta)] = f_{X,Y}(x, y)\delta^2$. 6.1 Conditional Distribution: $f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_{x,Y}(x,y)}$.
 - 6.2 Independence: $f_{X|Y}(x, y) = f_X(x)$

```
What is true? 
X has CDF F(x) and PDF f(x).
```

What is true? X has CDF F(x) and PDF f(x). (A) $Pr[X > t] = 1 - Pr[X \le t]$. (B) S(t) = Pr[X > t] = 1 - F(t). (C) Y = 2X, $f_Y(y) = 2f(y)$. (D) Y = 2X, $F_Y(y) = F(y/2)$. (E) Y = 2X, $f_Y(y) = \frac{1}{2}f(y/2)$.

What is true? X has CDF F(x) and PDF f(x). (A) $Pr[X > t] = 1 - Pr[X \le t]$. (B) S(t) = Pr[X > t] = 1 - F(t). (C) Y = 2X, $f_Y(y) = 2f(y)$. (D) Y = 2X, $F_Y(y) = F(y/2)$. (E) Y = 2X, $f_Y(y) = \frac{1}{2}f(y/2)$.

(A), (B), (D) think events, (E) think event and density.

What is true? X has CDF F(x) and PDF f(x). (A) $Pr[X > t] = 1 - Pr[X \le t]$. (B) S(t) = Pr[X > t] = 1 - F(t). (C) Y = 2X, $f_Y(y) = 2f(y)$. (D) Y = 2X, $F_Y(y) = F(y/2)$. (E) Y = 2X, $f_Y(y) = \frac{1}{2}f(y/2)$.

(A), (B), (D) think events, (E) think event and density.

(C) confuses probability density of outcome with value of outcome.

Discrete: Probability of outcome \rightarrow random variables, events.

Discrete: Probability of outcome \rightarrow random variables, events. Continuous: "outcome" is real number.

Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval.

Discrete: Probability of outcome \rightarrow random variables, events.

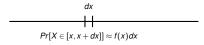
Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$

Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$

Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$



Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$

dx | | $Pr[X \in [x, x + dx]] \approx f(x)dx$

Joint Continuous in *d* variables: "outcome" is $\in \mathbb{R}^d$. Probability: Events is block.

Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$

 $\frac{dx}{||}$ $Pr[X \in [x, x + dx]] \approx f(x)dx$

Joint Continuous in *d* variables: "outcome" is $\in \mathbb{R}^d$. Probability: Events is block. Density: $Pr[(X, Y) \in ([x, x + dx], [y, y + dx])] = f(x, y)dxdy$

Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$

 $\frac{dx}{||}$ $Pr[X \in [x, x + dx]] \approx f(x)dx$

Joint Continuous in *d* variables: "outcome" is $\in \mathbb{R}^d$. Probability: Events is block. Density: $Pr[(X, Y) \in ([x, x + dx], [y, y + dx])] = f(x, y)dxdy$

Discrete: Probability of outcome \rightarrow random variables, events.

Continuous: "outcome" is real number. Probability: Events is interval. Density: $Pr[X \in [x, x + dx]] = f(x)dx$

dx $Pr[X \in [x, x + dx]] \approx f(x)dx$

dx

Joint Continuous in *d* variables: "outcome" is $\in \mathbb{R}^d$. Probability: Events is block. Density: $Pr[(X, Y) \in ([x, x + dx], [y, y + dx])] = f(x, y)dxdy$ $dy \longrightarrow Pr[(X, Y) \in ([x, x + dx], [y, y + dy])] \approx f(x, y)dxdy$

Probability!

Probability! Challenges us.

Probability! Challenges us. But really neat.

Probability! Challenges us. But really neat. At times,

Probability! Challenges us. But really neat. At times, continuous.

Probability! Challenges us. But really neat. At times, continuous. At others, discrete. Sample Space: Ω , $Pr[\omega]$.

Probability! Challenges us. But really neat. At times, continuous. At others, discrete. Sample Space: Ω , $Pr[\omega]$. Event: $Pr[A] = \sum_{\omega \in A} Pr[\omega]$ $\sum_{\omega} Pr[\omega] = 1$.

Probability! Challenges us. But really neat. At times, continuous. At others, discrete. Sample Space: Ω , $Pr[\omega]$. Event: $Pr[A] = \sum_{\omega \in A} Pr[\omega]$ $\sum_{\omega} Pr[\omega] = 1$.

Random Variable: XEvent: $A = [a, b], Pr[X \in A],$

Probability! Challenges us. But really neat. At times, continuous. At others, discrete. Sample Space: Ω , $Pr[\omega]$. Raute Pr[A] = $\sum_{\omega \in A} Pr[\omega]$ $\sum_{\omega} Pr[\omega] = 1$. Random variables: $X(\omega)$. Distribution: Pr[X = x] $\sum_{x} Pr[X = x] = 1$.

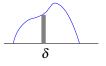
Random Variable: X Event: A = [a, b], $Pr[X \in A]$, CDF: $F(x) = Pr[X \le x]$. PDF: $f(x) = \frac{dF(x)}{dx}$. $\int_{-\infty}^{\infty} f(x) = 1$.

Probability! Challenges us. But really neat. At times, continuous. At others, discrete. Sample Space: Ω , $Pr[\omega]$. Raute Pr[A] = $\sum_{\omega \in A} Pr[\omega]$ $\sum_{\omega} Pr[\omega] = 1$. Random variables: $X(\omega)$. Distribution: Pr[X = x] $\sum_{x} Pr[X = x] = 1$.

Random Variable: X Event: A = [a, b], $Pr[X \in A]$, CDF: $F(x) = Pr[X \le x]$. PDF: $f(x) = \frac{dF(x)}{dx}$. $\int_{-\infty}^{\infty} f(x) = 1$.

Probability! Challenges us. But really neat. At times, continuous. At others, discrete. Sample Space: Ω , $Pr[\omega]$. Rauce $Pr[A] = \sum_{\omega \in A} Pr[\omega]$ $\sum_{\omega} Pr[\omega] = 1$. Random variables: $X(\omega)$. Distribution: Pr[X = x] $\sum_{x} Pr[X = x] = 1$.

Continuous as Discrete. $Pr[X \in [x, x + \delta]] \approx f(x)\delta$ Random Variable: X Event: $A = [a, b], Pr[X \in A],$ CDF: $F(x) = Pr[X \le x].$ PDF: $f(x) = \frac{dF(x)}{dx}.$ $\int_{-\infty}^{\infty} f(x) = 1.$



Conditional Probability.

Conditional Probability.

Events: A, B

Conditional Probability.

Events: A, B

Discrete: "Heads", "Tails", X = 1, Y = 5.

Conditional Probability.

Events: A, B

Discrete: "Heads", "Tails", X = 1, Y = 5.

Continuous: X in [.2,.3]. $X \in [.2,.3]$ or $X \in [.4,.6]$.

Conditional Probability. Events: A, BDiscrete: "Heads", "Tails", X = 1, Y = 5. Continuous: X in [.2,.3]. $X \in [.2,.3]$ or $X \in [.4,.6]$.

Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$

Conditional Probability. Events: *A*, *B* Discrete: "Heads", "Tails", X = 1, Y = 5. Continuous: *X* in [.2,.3]. $X \in [.2,.3]$ or $X \in [.4,.6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$ Pr["Second Heads"]"First Heads"],

 $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$

Conditional Probability. Events: A, BDiscrete: "Heads", "Tails", X = 1, Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3]|X \in [.2, .3]$ or $X \in [.5, .6]$].

Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$

Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]

Conditional Probability. Events: A.B. Discrete: "Heads", "Tails", X = 1, Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A\cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$ Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]

B is First coin heads.

Conditional Probability. Events: A.B. Discrete: "Heads", "Tails", X = 1, Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A\cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$ Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]B is First coin heads. $Pr[X \in [.45, .55]] = Pr[X \in [.45, .50]] + Pr[X \in (.50, .55]]$

Conditional Probability. Events: A.B. Discrete: "Heads", "Tails", X = 1, Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A\cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$ Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]B is First coin heads. $Pr[X \in [.45, .55]] = Pr[X \in [.45, .50]] + Pr[X \in (.50, .55]]$ *B* is $X \in [0, .5]$

Conditional Probability. Events: A.B. Discrete: "Heads", "Tails", X = 1, Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A\cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$ Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]B is First coin heads. $Pr[X \in [.45, .55]] = Pr[X \in [.45, .50]] + Pr[X \in (.50, .55]]$ *B* is $X \in [0, .5]$

Conditional Probability. Events: A.B. Discrete: "Heads", "Tails", X = 1. Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A\cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$ Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]B is First coin heads. $Pr[X \in [.45, .55]] = Pr[X \in [.45, .50]] + Pr[X \in (.50, .55]]$ *B* is $X \in [0, .5]$ Product Rule: $Pr[A \cap B] = Pr[A|B]Pr[B]$. Bayes Rule: Pr[A|B] = Pr[B|A]Pr[A]/Pr[B].

Conditional Probability. Events: A.B. Discrete: "Heads", "Tails", X = 1. Y = 5. Continuous: X in [.2, .3]. $X \in [.2, .3]$ or $X \in [.4, .6]$. Conditional Probability: $Pr[A|B] = \frac{Pr[A\cap B]}{Pr[B]}$ Pr["Second Heads"|"First Heads"], $Pr[X \in [.2, .3] | X \in [.2, .3] \text{ or } X \in [.5, .6]].$ Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$ Pr["Second Heads"] = Pr[HH] + Pr[HT]B is First coin heads. $Pr[X \in [.45, .55]] = Pr[X \in [.45, .50]] + Pr[X \in (.50, .55]]$ *B* is $X \in [0, .5]$ Product Rule: $Pr[A \cap B] = Pr[A|B]Pr[B]$. Bayes Rule: Pr[A|B] = Pr[B|A]Pr[A]/Pr[B].

All work for continuous with intervals as events.

Conditional Density: $f_{X|Y}(x, y)$.

Conditional Density: $f_{X|Y}(x, y)$.

Conditional Probability: $Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$

Conditional Density: $f_{X|Y}(x, y)$. Conditional Probability: $Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$ $Pr[X \in [x, x + dx] | Y \in [y, y + dy]] = \frac{f_{X,Y}(x, y) dx dy}{f_Y(y) dy}$

Conditional Density: $f_{X|Y}(x, y)$. Conditional Probability: $Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$ $Pr[X \in [x, x + dx] | Y \in [y, y + dy]] = \frac{f_{X,Y}(x,y)dxdy}{f_Y(y)dy}$ $f_{X|Y}(x, y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{f_{X,Y}(x,y)}{\int_{-\infty}^{+\infty} f_{X,Y}(x,y)dx}$

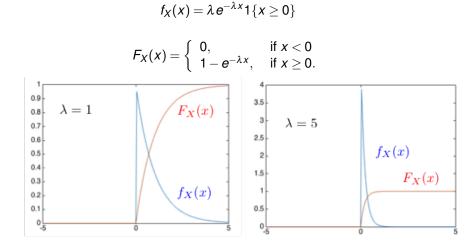
Conditional Density: $f_{X|Y}(x, y)$. Conditional Probability: $Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$ $Pr[X \in [x, x + dx] | Y \in [y, y + dy]] = \frac{f_{X,Y}(x,y)dxdy}{f_Y(y)dy}$ $f_{X|Y}(x, y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{f_{X,Y}(x,y)}{\int_{-\infty}^{+\infty} f_{X,Y}(x,y)dx}$

Corollary: For independent random variables, $f_{X|Y}(x, y) = f_X(x)$.

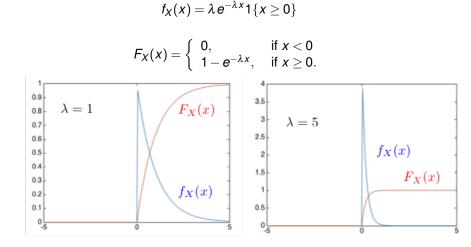
The exponential distribution with parameter $\lambda > 0$ is defined by

The exponential distribution with parameter $\lambda > 0$ is defined by $f_X(x) = \lambda e^{-\lambda x} \mathbf{1}\{x \ge 0\}$

The exponential distribution with parameter $\lambda > 0$ is defined by



The exponential distribution with parameter $\lambda > 0$ is defined by



Note that $Pr[X > t] = e^{-\lambda t}$ for t > 0.

1. Expo is memoryless.

1. *Expo* is memoryless. Let $X = Expo(\lambda)$.

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

 $Pr[X > t + s \mid X > s] =$

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} =$$

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling Expo.

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo.* Let $X = Expo(\lambda)$ and Y = aX for some a > 0.

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

Pr[Y > t] =

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

Pr[Y > t] = Pr[aX > t] =

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

$$Pr[Y > t] = Pr[aX > t] = Pr[X > t/a]$$

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

$$Pr[Y > t] = Pr[aX > t] = Pr[X > t/a]$$
$$= e^{-\lambda(t/a)} = e^{-(\lambda/a)t} =$$

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

$$\begin{aligned} \mathsf{Pr}[Y > t] &= \mathsf{Pr}[aX > t] = \mathsf{Pr}[X > t/a] \\ &= e^{-\lambda(t/a)} = e^{-(\lambda/a)t} = \mathsf{Pr}[Z > t] \text{ for } Z = \mathsf{Expo}(\lambda/a). \end{aligned}$$

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

$$\begin{aligned} \Pr[Y > t] &= \Pr[aX > t] = \Pr[X > t/a] \\ &= e^{-\lambda(t/a)} = e^{-(\lambda/a)t} = \Pr[Z > t] \text{ for } Z = Expo(\lambda/a). \end{aligned}$$

Thus, $a \times Expo(\lambda) = Expo(\lambda/a)$.

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

$$\begin{aligned} \Pr[Y > t] &= \Pr[aX > t] = \Pr[X > t/a] \\ &= e^{-\lambda(t/a)} = e^{-(\lambda/a)t} = \Pr[Z > t] \text{ for } Z = Expo(\lambda/a). \end{aligned}$$

Thus, $a \times Expo(\lambda) = Expo(\lambda/a)$. Also, $Expo(\lambda) = \frac{1}{\lambda} Expo(1)$.

3. Scaling Uniform.

3. Scaling Uniform. Let X = U[0, 1] and Y = a + bX where b > 0. Then,

 $Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] =$

3. Scaling Uniform. Let X = U[0, 1] and Y = a + bX where b > 0. Then,

 $Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] =$$

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for}$$

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y-a}{b} < 1$$

$$\begin{aligned} \Pr[Y \in (y, y + \delta)] &= \Pr[a + bX \in (y, y + \delta)] = \Pr[X \in (\frac{y - a}{b}, \frac{y + \delta - a}{b})] \\ &= \Pr[X \in (\frac{y - a}{b}, \frac{y - a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y - a}{b} < 1 \\ &= \frac{1}{b}\delta, \text{ for } a < y < a + b. \end{aligned}$$

3. Scaling Uniform. Let X = U[0, 1] and Y = a + bX where b > 0. Then,

$$\begin{aligned} \Pr[Y \in (y, y + \delta)] &= \Pr[a + bX \in (y, y + \delta)] = \Pr[X \in (\frac{y - a}{b}, \frac{y + \delta - a}{b})] \\ &= \Pr[X \in (\frac{y - a}{b}, \frac{y - a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y - a}{b} < 1 \\ &= \frac{1}{b}\delta, \text{ for } a < y < a + b. \end{aligned}$$

Thus, $f_Y(y) = \frac{1}{b}$ for a < y < a + b.

3. Scaling Uniform. Let X = U[0, 1] and Y = a + bX where b > 0. Then,

$$\begin{aligned} \Pr[Y \in (y, y + \delta)] &= \Pr[a + bX \in (y, y + \delta)] = \Pr[X \in (\frac{y - a}{b}, \frac{y + \delta - a}{b})] \\ &= \Pr[X \in (\frac{y - a}{b}, \frac{y - a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y - a}{b} < 1 \\ &= \frac{1}{b}\delta, \text{ for } a < y < a + b. \end{aligned}$$

Thus, $f_Y(y) = \frac{1}{b}$ for a < y < a+b. Hence, Y = U[a, a+b].

3. Scaling Uniform. Let X = U[0, 1] and Y = a + bX where b > 0. Then,

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y-a}{b} < 1$$
$$= \frac{1}{b}\delta, \text{ for } a < y < a+b.$$

Thus, $f_Y(y) = \frac{1}{b}$ for a < y < a + b. Hence, Y = U[a, a + b].

Replace b by b-a, use X = U[0,1], then Y = a + (b-a)X is U[a,b].

Some More Properties

4. Scaling pdf.

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0.

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

 $Pr[Y \in (y, y + \delta)] = Pr[a + bX \in (y, y + \delta)] =$

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

 $Pr[Y \in (y, y + \delta)] = Pr[a + bX \in (y, y + \delta)] = Pr[X \in (\frac{y - a}{b}, \frac{y + \delta - a}{b})]$

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] =$$

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = f_X(\frac{y-a}{b}, \frac{\delta}{b}).$$

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = f_X(\frac{y-a}{b})\frac{\delta}{b}.$$

Now, the left-hand side is

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = f_X(\frac{y-a}{b})\frac{\delta}{b}.$$

Now, the left-hand side is $f_Y(y)\delta$.

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = f_X(\frac{y-a}{b})\frac{\delta}{b}.$$

Now, the left-hand side is $f_Y(y)\delta$. Hence,

$$f_Y(y) = \frac{1}{b} f_X(\frac{y-a}{b})$$

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and Y = a + bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = f_X(\frac{y-a}{b})\frac{\delta}{b}.$$

Now, the left-hand side is $f_Y(y)\delta$. Hence,

$$f_Y(y) = \frac{1}{b} f_X(\frac{y-a}{b})$$



Definition: The expectation of a random variable X with pdf f(x) is *defined* as

Definition: The expectation of a random variable X with pdf f(x) is defined as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Definition: The expectation of a random variable X with pdf f(x) is defined as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification:

Definition: The expectation of a random variable X with pdf f(x) is *defined* as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$.

Definition: The expectation of a random variable X with pdf f(x) is *defined* as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

$$E[X] = \sum_{n} (n\delta) \Pr[X = n\delta]$$

Definition: The expectation of a random variable X with pdf f(x) is *defined* as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

$$E[X] = \sum_{n} (n\delta) Pr[X = n\delta] = \sum_{n} (n\delta) f_X(n\delta) \delta$$

Definition: The expectation of a random variable X with pdf f(x) is defined as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

$$E[X] = \sum_{n} (n\delta) Pr[X = n\delta] = \sum_{n} (n\delta) f_X(n\delta) \delta = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Definition: The expectation of a random variable X with pdf f(x) is defined as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

$$E[X] = \sum_{n} (n\delta) Pr[X = n\delta] = \sum_{n} (n\delta) f_X(n\delta) \delta = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Indeed, for any g, one has $\int g(x) dx \approx \sum_n g(n\delta) \delta$.

Definition: The expectation of a random variable X with pdf f(x) is defined as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

$$E[X] = \sum_{n} (n\delta) Pr[X = n\delta] = \sum_{n} (n\delta) f_X(n\delta) \delta = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Indeed, for any g, one has $\int g(x) dx \approx \sum_n g(n\delta)\delta$. Choose $g(x) = x f_X(x)$.

Definition: The expectation of a random variable X with pdf f(x) is defined as

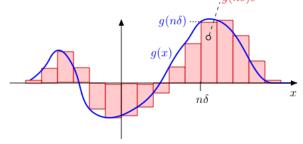
$$\Xi[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

1

$$E[X] = \sum_{n} (n\delta) Pr[X = n\delta] = \sum_{n} (n\delta) f_X(n\delta) \delta = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Indeed, for any *g*, one has $\int g(x) dx \approx \sum_n g(n\delta)\delta$. Choose $g(x) = x f_X(x)$.



1. X = U[0, 1].

1. X = U[0, 1]. Then, $f_X(x) =$

1. X = U[0,1]. Then, $f_X(x) = 1\{0 \le x \le 1\}$.

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

1. X = U[0, 1]. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x . 1 dx =$$

1. X = U[0, 1]. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 =$$

1. X = U[0, 1]. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

2. X = distance to 0 of dart shot uniformly in unit circle.

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 2x dx =$$

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 2x dx = \left[\frac{2x^3}{3}\right]_0^1 =$$

1.
$$X = U[0, 1]$$
. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 1 dx = \left[\frac{x^2}{2}\right]_0^1 = \frac{1}{2}.$$

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 2x dx = \left[\frac{2x^3}{3}\right]_0^1 = \frac{2}{3}.$$

3. $X = Expo(\lambda)$.

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$.

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\int_a^b u(x)dv(x) = \left[u(x)v(x)\right]_a^b - \int_a^b v(x)du(x)$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\int_a^b u(x)dv(x) = [u(x)v(x)]_a^b - \int_a^b v(x)du(x)$$
$$= u(b)v(b) - u(a)v(a) - \int_a^b v(x)du(x).$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\int_a^b u(x)dv(x) = [u(x)v(x)]_a^b - \int_a^b v(x)du(x)$$
$$= u(b)v(b) - u(a)v(a) - \int_a^b v(x)du(x).$$

Thus,

$$\int_0^\infty x de^{-\lambda x} = [xe^{-\lambda x}]_0^\infty - \int_0^\infty e^{-\lambda x} dx$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\int_a^b u(x)dv(x) = [u(x)v(x)]_a^b - \int_a^b v(x)du(x)$$
$$= u(b)v(b) - u(a)v(a) - \int_a^b v(x)du(x).$$

Thus,

$$\int_0^\infty x de^{-\lambda x} = [xe^{-\lambda x}]_0^\infty - \int_0^\infty e^{-\lambda x} dx$$
$$= 0 - 0 + \frac{1}{\lambda} \int_0^\infty de^{-\lambda x} =$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\int_a^b u(x)dv(x) = [u(x)v(x)]_a^b - \int_a^b v(x)du(x)$$
$$= u(b)v(b) - u(a)v(a) - \int_a^b v(x)du(x).$$

Thus,

$$\int_0^\infty x de^{-\lambda x} = [xe^{-\lambda x}]_0^\infty - \int_0^\infty e^{-\lambda x} dx$$
$$= 0 - 0 + \frac{1}{\lambda} \int_0^\infty de^{-\lambda x} = -\frac{1}{\lambda}$$

3.
$$X = Expo(\lambda)$$
. Then, $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$. Thus,

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\int_a^b u(x)dv(x) = [u(x)v(x)]_a^b - \int_a^b v(x)du(x)$$
$$= u(b)v(b) - u(a)v(a) - \int_a^b v(x)du(x).$$

Thus,

$$\int_0^\infty x de^{-\lambda x} = [xe^{-\lambda x}]_0^\infty - \int_0^\infty e^{-\lambda x} dx$$
$$= 0 - 0 + \frac{1}{\lambda} \int_0^\infty de^{-\lambda x} = -\frac{1}{\lambda}.$$

Hence, $E[X] = \frac{1}{\lambda}$.

For any μ and σ , a **normal** (aka **Gaussian**)

For any μ and σ , a **normal** (aka **Gaussian**) random variable *Y*, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

For any μ and σ , a **normal** (aka **Gaussian**) random variable *Y*, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/2\sigma^2}$$

For any μ and σ , a **normal** (aka **Gaussian**) random variable *Y*, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

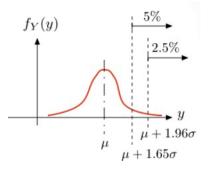
$$f_Y(y) = rac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/2\sigma^2}$$

Standard normal has $\mu = 0$ and $\sigma = 1$.

For any μ and σ , a **normal** (aka **Gaussian**) random variable *Y*, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

$$f_Y(y) = rac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/2\sigma^2}$$

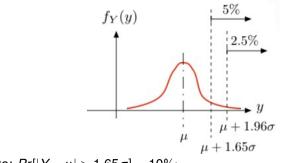
Standard normal has $\mu = 0$ and $\sigma = 1$.



For any μ and σ , a **normal** (aka **Gaussian**) random variable *Y*, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

$$f_Y(y) = rac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/2\sigma^2}$$

Standard normal has $\mu = 0$ and $\sigma = 1$.

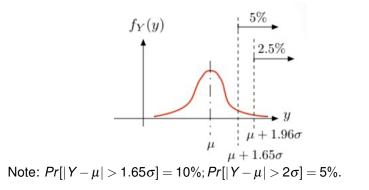


Note: $Pr[|Y - \mu| > 1.65\sigma] = 10\%$;

For any μ and σ , a **normal** (aka **Gaussian**) random variable *Y*, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

$$f_Y(y) = rac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/2\sigma^2}$$

Standard normal has $\mu = 0$ and $\sigma = 1$.



Scaling and Shifting and properties

Theorem Let $X = \mathcal{N}(0, 1)$ and $Y = \mu + \sigma X$. Then

 $Y = \mathcal{N}(\mu, \sigma^2).$

Scaling and Shifting and properties

Theorem Let $X = \mathcal{N}(0, 1)$ and $Y = \mu + \sigma X$. Then

$$Y = \mathcal{N}(\mu, \sigma^2).$$

Theorem If $Y = \mathcal{N}(\mu, \sigma^2)$, then $E[Y] = \mu$

Scaling and Shifting and properties

Theorem Let $X = \mathcal{N}(0, 1)$ and $Y = \mu + \sigma X$. Then

$$Y = \mathcal{N}(\mu, \sigma^2).$$

Theorem If $Y = \mathcal{N}(\mu, \sigma^2)$, then

$$E[Y] = \mu$$
 and $var[Y] = \sigma^2$.

Theorem: Set of independent identically distributed random variables, X_i , $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Theorem: Set of independent identically distributed random variables, X_i , $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Say X_i have expectation $\mu = E(X_i)$ and variance σ^2 .

Theorem: Set of independent identically distributed random variables, X_i ,

 $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Say X_i have expectation $\mu = E(X_i)$ and variance σ^2 .

Mean of A_n is μ , and variance is σ^2/n .

Theorem: Set of independent identically distributed random variables, X_i ,

 $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Say X_i have expectation $\mu = E(X_i)$ and variance σ^2 .

Mean of A_n is μ , and variance is σ^2/n .

$$Pr[|A_n - \mu| > \varepsilon] \le$$

Theorem: Set of independent identically distributed random variables, X_i ,

 $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Say X_i have expectation $\mu = E(X_i)$ and variance σ^2 .

Mean of A_n is μ , and variance is σ^2/n .

$$Pr[|A_n - \mu| > \varepsilon] \le rac{var[A_n]}{\varepsilon^2} =$$

Theorem: Set of independent identically distributed random variables, X_i ,

 $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Say X_i have expectation $\mu = E(X_i)$ and variance σ^2 .

Mean of A_n is μ , and variance is σ^2/n .

$$\Pr[|A_n - \mu| > \varepsilon] \le rac{var[A_n]}{\varepsilon^2} = rac{\sigma^2}{n\varepsilon}$$

Theorem: Set of independent identically distributed random variables, X_i ,

 $A_n = \frac{1}{n} \sum X_i$ "tends to the mean."

Say X_i have expectation $\mu = E(X_i)$ and variance σ^2 .

Mean of A_n is μ , and variance is σ^2/n .

$$\Pr[|A_n - \mu| > \varepsilon] \leq rac{var[A_n]}{\varepsilon^2} = rac{\sigma^2}{n\varepsilon} o 0.$$

Central Limit Theorem

Central Limit Theorem

Let X_1, X_2, \ldots be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$.

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n \rightarrow \mathcal{N}(0,1), \text{as } n \rightarrow \infty.$$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n
ightarrow \mathcal{N}(0,1),$$
 as $n
ightarrow \infty$.

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n \rightarrow \mathcal{N}(0,1), \text{as } n \rightarrow \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof:

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n
ightarrow \mathcal{N}(0,1),$$
 as $n
ightarrow \infty$.

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n \rightarrow \mathcal{N}(0,1), \text{as } n \rightarrow \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n \to \mathcal{N}(0,1), \text{as } n \to \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

$$E(S_n)$$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n \rightarrow \mathcal{N}(0,1), \text{as } n \rightarrow \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

$$E(S_n) = \frac{1}{\sigma/\sqrt{n}}(E(A_n) - \mu)$$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n
ightarrow \mathcal{N}(0,1),$$
 as $n
ightarrow \infty$.

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

$$E(S_n) = \frac{1}{\sigma/\sqrt{n}}(E(A_n) - \mu) = 0$$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n \rightarrow \mathcal{N}(0,1), \text{as } n \rightarrow \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

Note:

$$E(S_n) = \frac{1}{\sigma/\sqrt{n}}(E(A_n) - \mu) = 0$$

 $Var(S_n)$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n
ightarrow \mathcal{N}(0,1), ext{as } n
ightarrow \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

$$E(S_n) = \frac{1}{\sigma/\sqrt{n}}(E(A_n) - \mu) = 0$$
$$Var(S_n) = \frac{1}{\sigma^2/n} Var(A_n)$$

Central Limit Theorem

Let $X_1, X_2, ...$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define $S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}.$

Then,

$$S_n
ightarrow \mathcal{N}(0,1), ext{as } n
ightarrow \infty.$$

That is,

$$Pr[S_n \leq \alpha] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

$$E(S_n) = \frac{1}{\sigma/\sqrt{n}}(E(A_n) - \mu) = 0$$
$$Var(S_n) = \frac{1}{\sigma^2/n} Var(A_n) = 1.$$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence $1 - \delta$ we need

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence $1 - \delta$ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence 1 – δ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

 $Pr[|A_n - \mu| > \varepsilon] \le$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence 1 – δ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

$$Pr[|A_n - \mu| > \varepsilon] \le C \int_{x \ge \varepsilon}^{\infty} e^{-rac{x^2}{2 varA}} \le C e^{-rac{\varepsilon^2}{2 varA}}$$

for $\varepsilon > \sqrt{VarA}$ (*C* is roughly $2/\sqrt{2\pi}$)

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence 1 – δ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

$$Pr[|A_n - \mu| > \varepsilon] \le C \int_{x \ge \varepsilon}^{\infty} e^{-rac{x^2}{2 varA}} \le C e^{-rac{\varepsilon^2}{2 varA}}$$

for $\varepsilon > \sqrt{VarA}$ (*C* is roughly $2/\sqrt{2\pi}$) Implies to get confidence $1 - C\delta$ we need

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence 1 – δ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

$$\Pr[|A_n - \mu| > \varepsilon] \le C \int_{x \ge \varepsilon}^{\infty} e^{-\frac{x^2}{2 \operatorname{var} A}} \le C e^{-\frac{\varepsilon^2}{2 \operatorname{var} A}}$$

for $\varepsilon > \sqrt{VarA}$ (*C* is roughly $2/\sqrt{2\pi}$)

Implies to get confidence $1 - C\delta$ we need

$$e^{-rac{arepsilon^2}{2 extsf{varA}}} \leq \delta$$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence 1 – δ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

$$\Pr[|A_n - \mu| > \varepsilon] \le C \int_{x \ge \varepsilon}^{\infty} e^{-\frac{x^2}{2 \operatorname{var} A}} \le C e^{-\frac{\varepsilon^2}{2 \operatorname{var} A}}$$

for $\varepsilon > \sqrt{VarA}$ (*C* is roughly $2/\sqrt{2\pi}$)

Implies to get confidence $1 - C\delta$ we need

$$e^{-rac{arepsilon^2}{2 ext{varA}}} \leq \delta \implies -rac{narepsilon^2}{2\sigma^2} \leq \log \delta$$

Recall: $A_n = \frac{1}{n} \sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2 . Mean of A_n is μ , and variance is σ^2/n .

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \le \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence 1 – δ we need

$$varA_n\varepsilon^2 = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \le \delta \text{ or } n \ge \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

$$\Pr[|A_n - \mu| > \varepsilon] \le C \int_{x \ge \varepsilon}^{\infty} e^{-\frac{x^2}{2 \operatorname{var} A}} \le C e^{-\frac{\varepsilon^2}{2 \operatorname{var} A}}$$

for $\varepsilon > \sqrt{VarA}$ (*C* is roughly $2/\sqrt{2\pi}$)

Implies to get confidence $1 - C\delta$ we need

$$e^{-rac{arepsilon^2}{2 ext{varA}}} \leq \delta \implies -rac{narepsilon^2}{2\sigma^2} \leq \log\delta \implies n \geq rac{2\sigma^2}{arepsilon^2}\lograc{1}{\delta}.$$

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

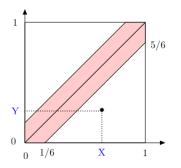
They agree they will wait for 10 minutes.

What is the probability they meet?

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

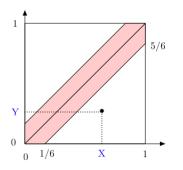
What is the probability they meet?



Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?

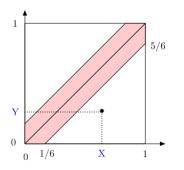


Here, (X, Y) are the times when the friends reach the restaurant.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



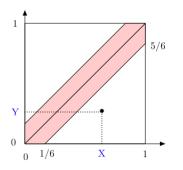
Here, (X, Y) are the times when the friends reach the restaurant.

The shaded area are the pairs where |X - Y| < 1/6,

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



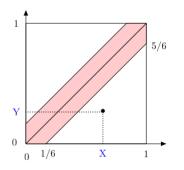
Here, (X, Y) are the times when the friends reach the restaurant.

The shaded area are the pairs where |X - Y| < 1/6, i.e., such that they meet.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



Here, (X, Y) are the times when the friends reach the restaurant.

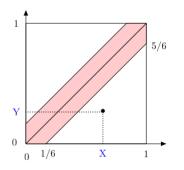
The shaded area are the pairs where |X - Y| < 1/6, i.e., such that they meet.

The complement is the sum of two rectangles.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



Here, (X, Y) are the times when the friends reach the restaurant.

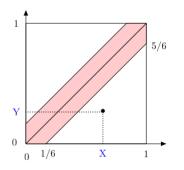
The shaded area are the pairs where |X - Y| < 1/6, i.e., such that they meet.

The complement is the sum of two rectangles. When you put them together, they form a square

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



Here, (X, Y) are the times when the friends reach the restaurant.

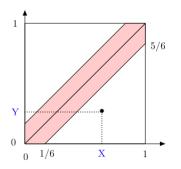
The shaded area are the pairs where |X - Y| < 1/6, i.e., such that they meet.

The complement is the sum of two rectangles. When you put them together, they form a square with sides 5/6.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



Thus, $Pr[meet] = 1 - (\frac{5}{6})^2 =$

Here, (X, Y) are the times when the friends reach the restaurant.

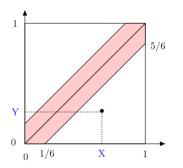
The shaded area are the pairs where |X - Y| < 1/6, i.e., such that they meet.

The complement is the sum of two rectangles. When you put them together, they form a square with sides 5/6.

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



Thus, $Pr[meet] = 1 - (\frac{5}{6})^2 = \frac{11}{36}$.

Here, (X, Y) are the times when the friends reach the restaurant.

The shaded area are the pairs where |X - Y| < 1/6, i.e., such that they meet.

The complement is the sum of two rectangles. When you put them together, they form a square with sides 5/6.

Breaking a Stick

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

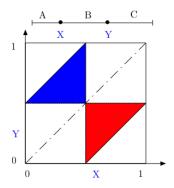
Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?

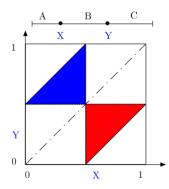
You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



You break a stick at two points chosen independently uniformly at random.

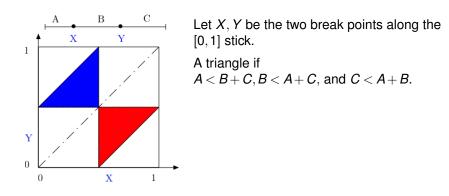
What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the [0, 1] stick.

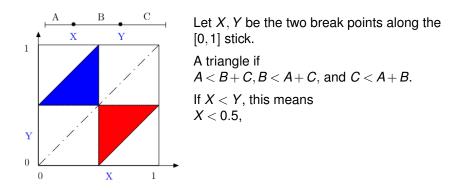
You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



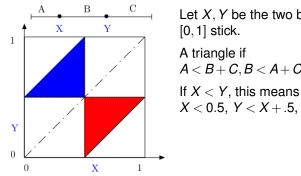
You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?

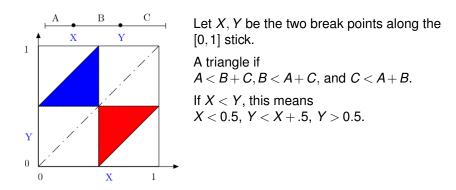


Let X, Y be the two break points along the

A < B + C, B < A + C, and C < A + B.

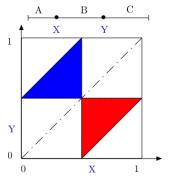
You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



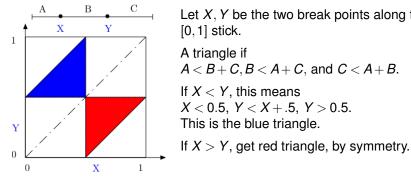
Let X, Y be the two break points along the [0, 1] stick.

A triangle if A < B + C, B < A + C, and C < A + B.

If X < Y, this means X < 0.5, Y < X + .5, Y > 0.5. This is the blue triangle.

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?

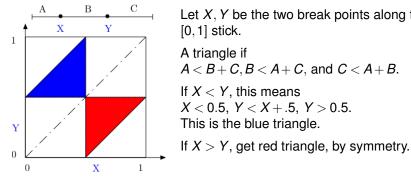


Let X, Y be the two break points along the

A < B + C, B < A + C, and C < A + B.

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?

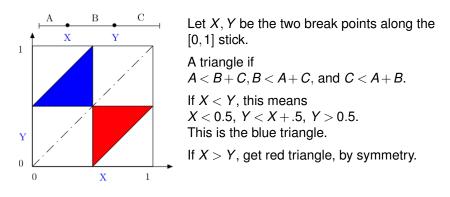


Let X, Y be the two break points along the

A < B + C, B < A + C, and C < A + B.

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Thus, Pr[make triangle] = 1/4.

Let $X = Expo(\lambda)$ and $Y = Expo(\mu)$ be independent.

Let $X = Expo(\lambda)$ and $Y = Expo(\mu)$ be independent. Define $Z = \max\{X, Y\}$.

Let $X = Expo(\lambda)$ and $Y = Expo(\mu)$ be independent. Define $Z = \max\{X, Y\}$.

Calculate E[Z].

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max{\{X, Y\}}$.

Calculate E[Z].

We compute f_Z , then integrate.

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max{X, Y}$.

Calculate E[Z].

We compute f_Z , then integrate.

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max\{X, Y\}$.

Calculate E[Z].

We compute f_Z , then integrate.

One has

Pr[Z < z] = Pr[X < z, Y < z]

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max{X, Y}$.

Calculate E[Z].

We compute f_Z , then integrate.

$$Pr[Z < z] = Pr[X < z, Y < z] = Pr[X < z]Pr[Y < z]$$

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max{X, Y}$.

Calculate E[Z].

We compute f_Z , then integrate.

$$Pr[Z < z] = Pr[X < z, Y < z] = Pr[X < z]Pr[Y < z]$$

= $(1 - e^{-\lambda z})(1 - e^{-\mu z}) =$

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max\{X, Y\}$.

Calculate E[Z].

We compute f_Z , then integrate.

$$Pr[Z < z] = Pr[X < z, Y < z] = Pr[X < z]Pr[Y < z]$$

= $(1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda + \mu)z}$

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max{X, Y}$.

Calculate E[Z].

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z] \Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda + \mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu) e^{-(\lambda + \mu)z}, \forall z > 0.$$

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max\{X, Y\}$.

Calculate E[Z].

We compute f_Z , then integrate.

One has

$$Pr[Z < z] = Pr[X < z, Y < z] = Pr[X < z]Pr[Y < z]$$

= $(1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda + \mu)z}$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu) e^{-(\lambda + \mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x \lambda e^{-\lambda x} dx = \lambda \left[-\frac{xe^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2} \right]_0^\infty = \frac{1}{\lambda}.$

$$E[Z] = \int_0^\infty z f_Z(z) dz =$$

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max\{X, Y\}$.

Calculate E[Z].

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z] \Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda + \mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu) e^{-(\lambda + \mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x \lambda e^{-\lambda x} dx = \lambda \left[-\frac{xe^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2} \right]_0^\infty = \frac{1}{\lambda}$.

$$E[Z] = \int_0^\infty z f_Z(z) dz = \frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}.$$

Let
$$X = Expo(\lambda)$$
 and $Y = Expo(\mu)$ be independent.
Define $Z = \max\{X, Y\}$.

Calculate E[Z].

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z] \Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda + \mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu) e^{-(\lambda + \mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x \lambda e^{-\lambda x} dx = \lambda \left[-\frac{xe^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2} \right]_0^\infty = \frac{1}{\lambda}$.

$$E[Z] = \int_0^\infty z f_Z(z) dz = \frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}.$$

Let X_1, \ldots, X_n be i.i.d. Expo(1).

Let $X_1, ..., X_n$ be i.i.d. *Expo*(1). Define $Z = \min\{X_1, X_2, ..., X_n\}$.

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \min\{X_1, X_2, \ldots, X_n\}$. What is true?

Let $X_1, ..., X_n$ be i.i.d. Expo(1). Define $Z = \min\{X_1, X_2, ..., X_n\}$. What is true?

(A) Z is exponential. (B) Parameter is n. (C) $\lim_{N\to\infty} (1-n/N)^N \to e^{-n}$ (D) E[Z] = 1/n.

Let $X_1, ..., X_n$ be i.i.d. Expo(1). Define $Z = \min\{X_1, X_2, ..., X_n\}$. What is true?

(A) Z is exponential. (B) Parameter is n. (C) $\lim_{N\to\infty} (1-n/N)^N \to e^{-n}$ (D) E[Z] = 1/n.

(C) is an argument for (A), (B) and (D).

Let X_1, \ldots, X_n be i.i.d. Expo(1).

Let $X_1, ..., X_n$ be i.i.d. *Expo*(1). Define $Z = \max\{X_1, X_2, ..., X_n\}$.

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion.

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

From memoryless property of the exponential.

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$.

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$A_n = E[\min\{X_1,\ldots,X_n\}] + A_{n-1}$$

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$A_n = E[\min\{X_1, \dots, X_n\}] + A_{n-1}$$

= $\frac{1}{n} + A_{n-1}$

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$A_n = E[\min\{X_1, \dots, X_n\}] + A_{n-1}$$

= $\frac{1}{n} + A_{n-1}$

because the minimum of *Expo* is *Expo* with the sum of the rates.

Let X_1, \ldots, X_n be i.i.d. Expo(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$. Calculate E[Z].

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}.$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$A_n = E[\min\{X_1, \dots, X_n\}] + A_{n-1}$$

= $\frac{1}{n} + A_{n-1}$

because the minimum of *Expo* is *Expo* with the sum of the rates. Hence,

$$E[Z] = A_n = 1 + \frac{1}{2} + \dots + \frac{1}{n} = H(n).$$

In digital video and audio, one represents a continuous value by a finite number of bits.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model:

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

Analysis:

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

Analysis: We see that *Z* is uniform in $[0, a = 2^{-(n+1)}]$.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

Analysis: We see that *Z* is uniform in $[0, a = 2^{-(n+1)}]$.

Thus,

$$E[Z^2] = \frac{a^2}{3} = \frac{1}{3}2^{-2(n+1)}.$$

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

Analysis: We see that *Z* is uniform in $[0, a = 2^{-(n+1)}]$.

Thus,

$$E[Z^2] = \frac{a^2}{3} = \frac{1}{3}2^{-2(n+1)}.$$

The power of the signal X is $E[X^2] =$

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

Analysis: We see that *Z* is uniform in $[0, a = 2^{-(n+1)}]$.

Thus,

$$E[Z^2] = \frac{a^2}{3} = \frac{1}{3}2^{-2(n+1)}.$$

The power of the signal X is $E[X^2] = \frac{1}{3}$.

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: X = U[0, 1] is the continuous value. *Y* is the closest multiple of 2^{-n} to *X*. Thus, we can represent *Y* with *n* bits. The error is Z := X - Y.

The power of the noise is $E[Z^2]$.

Analysis: We see that *Z* is uniform in $[0, a = 2^{-(n+1)}]$.

Thus,

$$E[Z^2] = \frac{a^2}{3} = \frac{1}{3}2^{-2(n+1)}.$$

The power of the signal X is $E[X^2] = \frac{1}{3}$.

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$. The signal to noise ratio (SNR)

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

Thus,

$$SNR = 2^{2(n+1)}$$
.

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

Thus,

$$SNR = 2^{2(n+1)}$$
.

Expressed in decibels, one has

 $SNR(dB) = 10 \log_{10}(SNR)$

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

Thus,

$$SNR = 2^{2(n+1)}$$
.

Expressed in decibels, one has

 $SNR(dB) = 10 \log_{10}(SNR) = 20(n+1) \log_{10}(2)$

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

Thus,

$$SNR = 2^{2(n+1)}$$
.

Expressed in decibels, one has

 $SNR(dB) = 10 \log_{10}(SNR) = 20(n+1) \log_{10}(2) \approx 6(n+1).$

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

Thus,

$$SNR = 2^{2(n+1)}$$
.

Expressed in decibels, one has

```
SNR(dB) = 10 \log_{10}(SNR) = 20(n+1) \log_{10}(2) \approx 6(n+1).
```

For instance, if n = 16, then $SNR(dB) \approx 112 dB$.

Problem 1:

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis:

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

$$E[(X - Y)^2] =$$

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

$$E[(X-Y)^2] = E[X^2+Y^2-2XY]$$

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$
$$= \frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$$

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2:

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

Analysis:

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

$$E[||\mathbf{X} - \mathbf{Y}||^2] =$$

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

$$E[||\mathbf{X} - \mathbf{Y}||^2] = E[(X_1 - Y_1)^2] + E[(X_2 - Y_2)^2]$$

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

$$E[||\mathbf{X} - \mathbf{Y}||^2] = E[(X_1 - Y_1)^2] + E[(X_2 - Y_2)^2]$$

= $2 \times \frac{1}{6}$.

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

Analysis: One has

$$E[||\mathbf{X} - \mathbf{Y}||^2] = E[(X_1 - Y_1)^2] + E[(X_2 - Y_2)^2]$$

= $2 \times \frac{1}{6}$.

Problem 3:

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

Analysis: One has

$$E[||\mathbf{X} - \mathbf{Y}||^2] = E[(X_1 - Y_1)^2] + E[(X_2 - Y_2)^2]$$

= $2 \times \frac{1}{6}$.

Problem 3: What about in n dimensions?

Problem 1: Pick two points X and Y independently and uniformly at random in [0, 1].

What is $E[(X - Y)^2]$?

Analysis: One has

$$E[(X - Y)^{2}] = E[X^{2} + Y^{2} - 2XY]$$

= $\frac{1}{3} + \frac{1}{3} - 2\frac{1}{2}\frac{1}{2}$
= $\frac{2}{3} - \frac{1}{2} = \frac{1}{6}.$

Problem 2: What about in a unit square?

Analysis: One has

$$E[||\mathbf{X} - \mathbf{Y}||^2] = E[(X_1 - Y_1)^2] + E[(X_2 - Y_2)^2]$$

= $2 \times \frac{1}{6}$.

Problem 3: What about in *n* dimensions? $\frac{n}{6}$.







Continuous Probability

Continuous RVs are similar to discrete RVs

- Continuous RVs are similar to discrete RVs
- ▶ Think that $X \in [x, x + \varepsilon]$ with probability $f_X(x)\varepsilon$

- Continuous RVs are similar to discrete RVs
- ▶ Think that $X \in [x, x + \varepsilon]$ with probability $f_X(x)\varepsilon$
- Sums become integrals,

- Continuous RVs are similar to discrete RVs
- ▶ Think that $X \in [x, x + \varepsilon]$ with probability $f_X(x)\varepsilon$
- Sums become integrals,
- The exponential distribution is magical:

- Continuous RVs are similar to discrete RVs
- ▶ Think that $X \in [x, x + \varepsilon]$ with probability $f_X(x)\varepsilon$
- Sums become integrals,
- The exponential distribution is magical: memoryless.